

# Bayesian Model Averaging of Longitudinal Dose-Response Models

*Richard Payne, PhD*

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# Collaborators

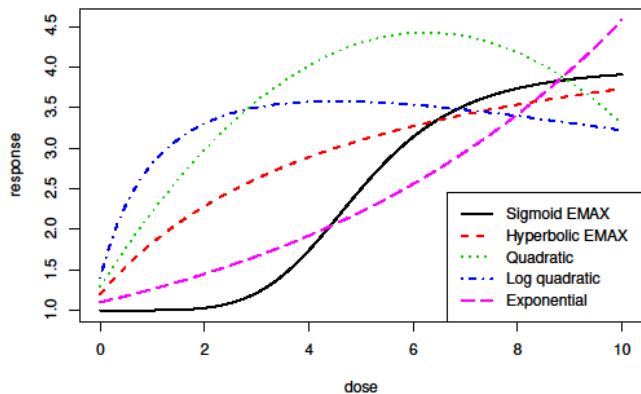
- Pallavi Ray
- Mitch Thomann

# Dose Response Modeling

- Why is it important?
  - 15% of failed first-time applications for New Molecular Entities to the FDA were related to uncertainties in dose selection (Sacks 2014).
  - Increasing the number of doses may not improve power, but it provides much greater information on the dose-response curve.
  - Helpful when choosing and/or justifying a dose to regulatory bodies.

# Dose Response Models

- What are they?
  - A model which assumes a (non-)parametric form across the dose range, so information across all arms is shared for estimation.
  - Different than pairwise tests which compare doses independently.
  - Examples:
    - EMAX (sigmoid and hyperbolic),
    - Quadratic, Exponential, Beta,
    - Linear, Log-linear, Log-quadratic...



With limited data, how does one choose a dose-response model *a priori*?

# Choosing a Dose-Response Model

- If there is prior data or information about the dose-response curve (e.g. other molecules in the same class and/or indication) or something is known about the dose-response curve (e.g. monotonic).
  - A suitable dose-response model might be able to be chosen *a priori*.
- How many doses? Two, three, more?
  - More flexible models will require more doses for suitable estimation (e.g. EMAX).
- What if there's not much data? What if there's a plausible scientific hypothesis of non-monotonicity, but it's unclear at which dose this might occur?
  - This is a good candidate for using Bayesian model averaging.

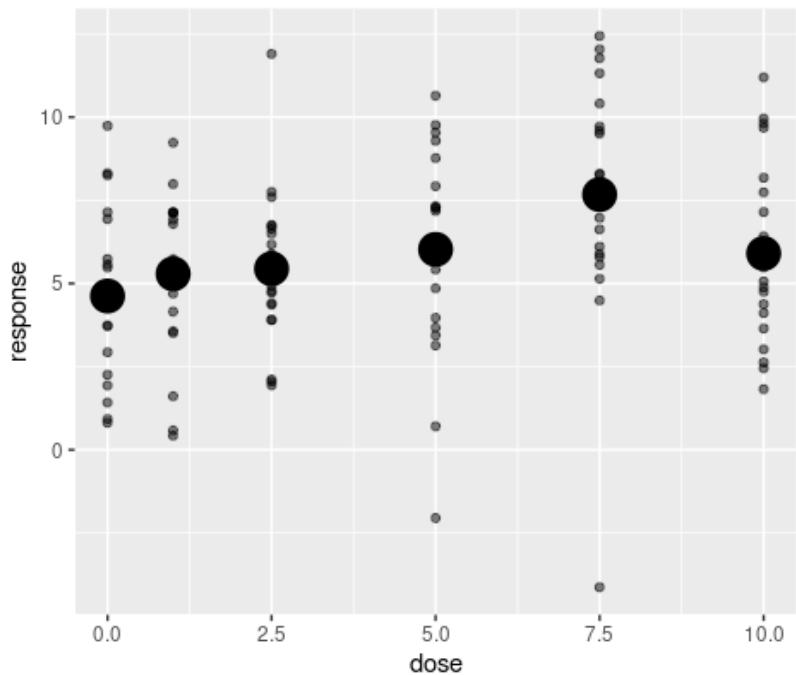
# Hypothetical Scenario

- Suppose there's a new molecule with limited phase I data.
- There is no data from other molecules in the same class.
- There is a plausible scientific hypothesis that the dose-response is non-monotonic at high enough doses, but the dose range proposed in the phase 2 dose-finding study is not believed to be in that zone.
- To account for the possibility of non-monotonicity, Bayesian model averaging is used with prior weight of .75 on an EMAX model, and .25 on a quadratic model.

# Example

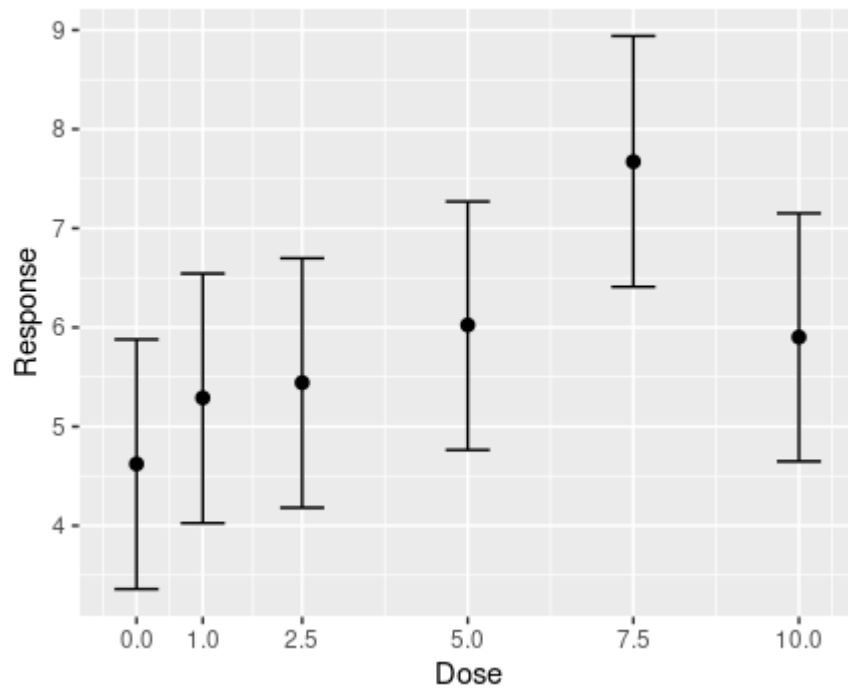
Pairwise p-value (vs placebo)  
is significant for dose 7.5  
(0.005) but not dose 10 (0.14).

## Simulated Data



## Independent Bayesian Credible intervals

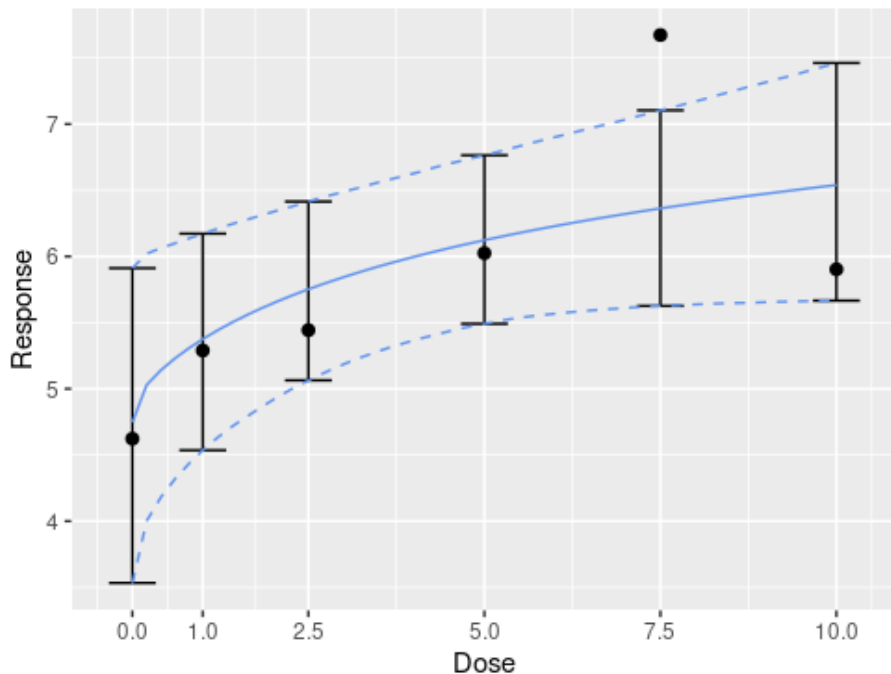
2.5%, 97.5% Posterior Quantiles



# Example: Dose Response Models

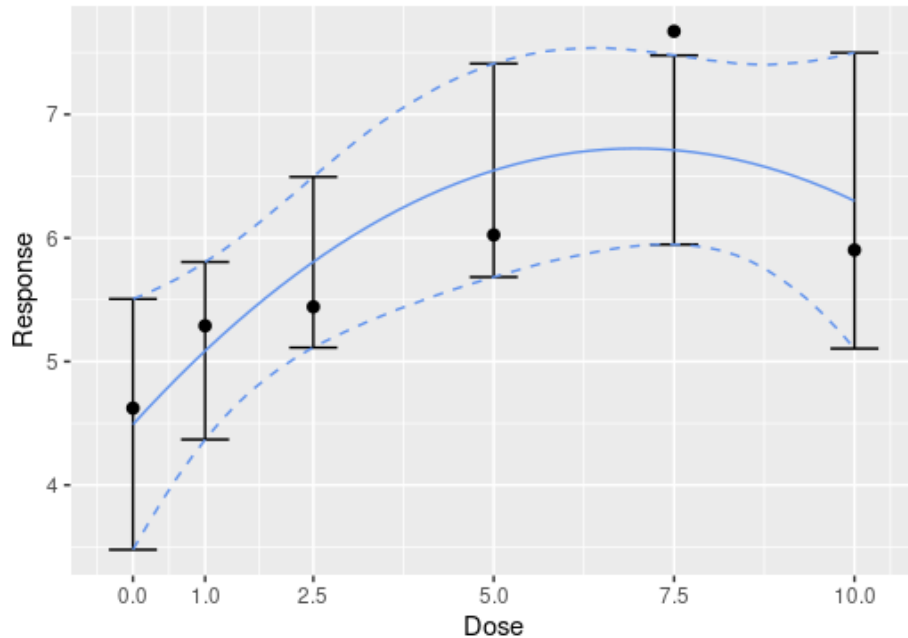
## EMAX Fit

Posterior mean (solid) and 2.5%, 97.5% quantiles (dashed)



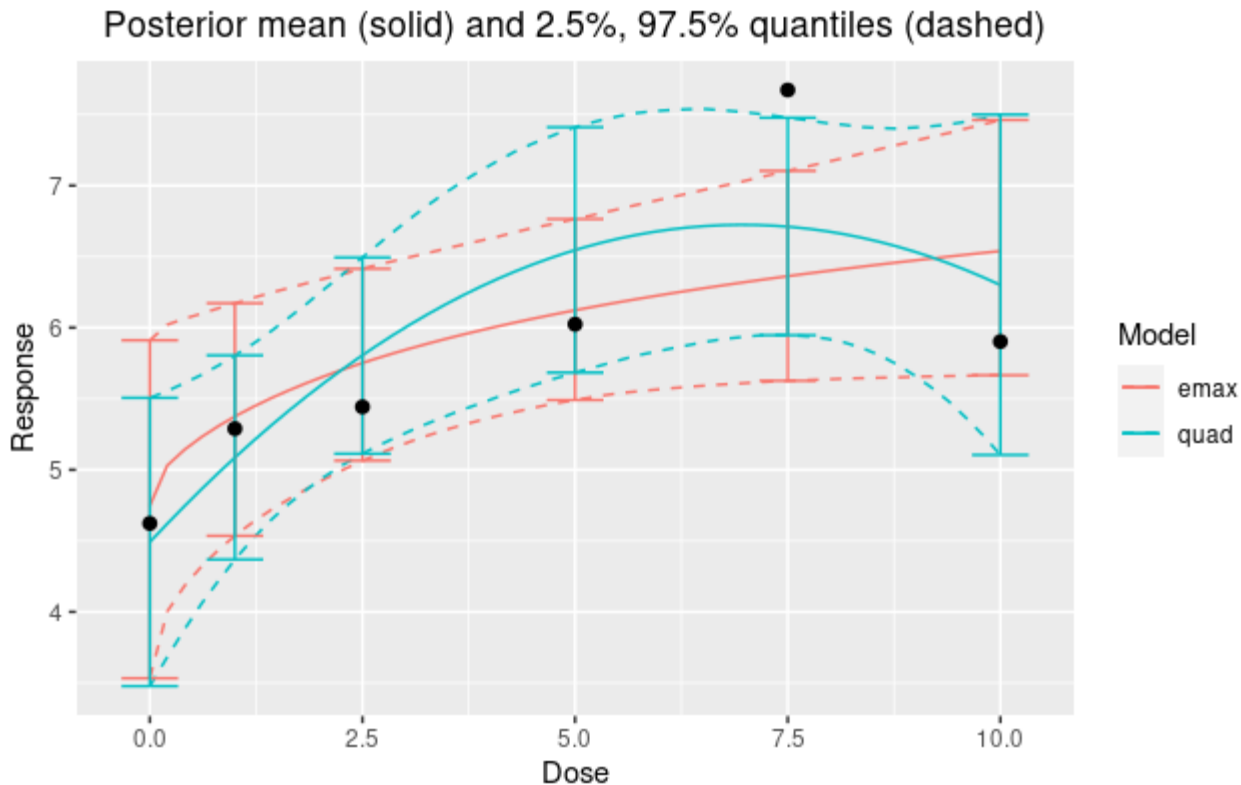
## Quadratic Fit

Posterior mean (solid) and 2.5%, 97.5% quantiles (dashed)





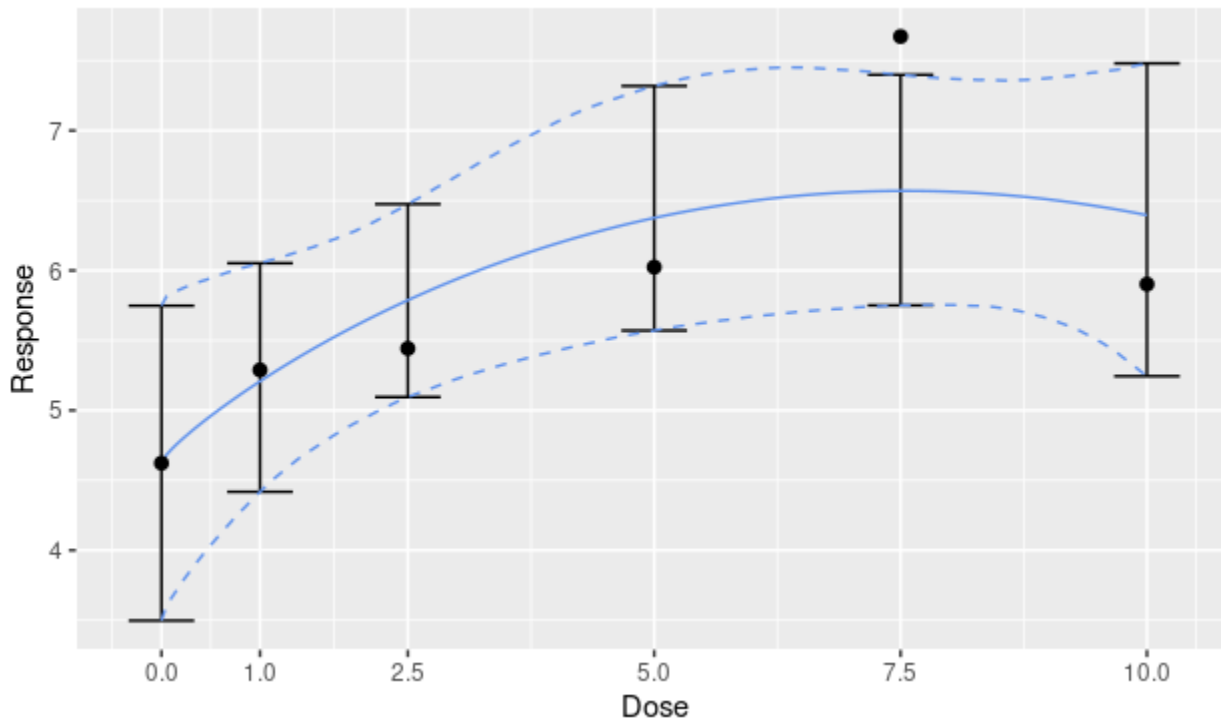
# Example: Dose Response Models



# Example: Bayesian Model Averaging

## EMAX & Quadratic Models

Posterior mean (solid) and 2.5%, 97.5% quantiles (dashed)



# Bayesian Model Averaging

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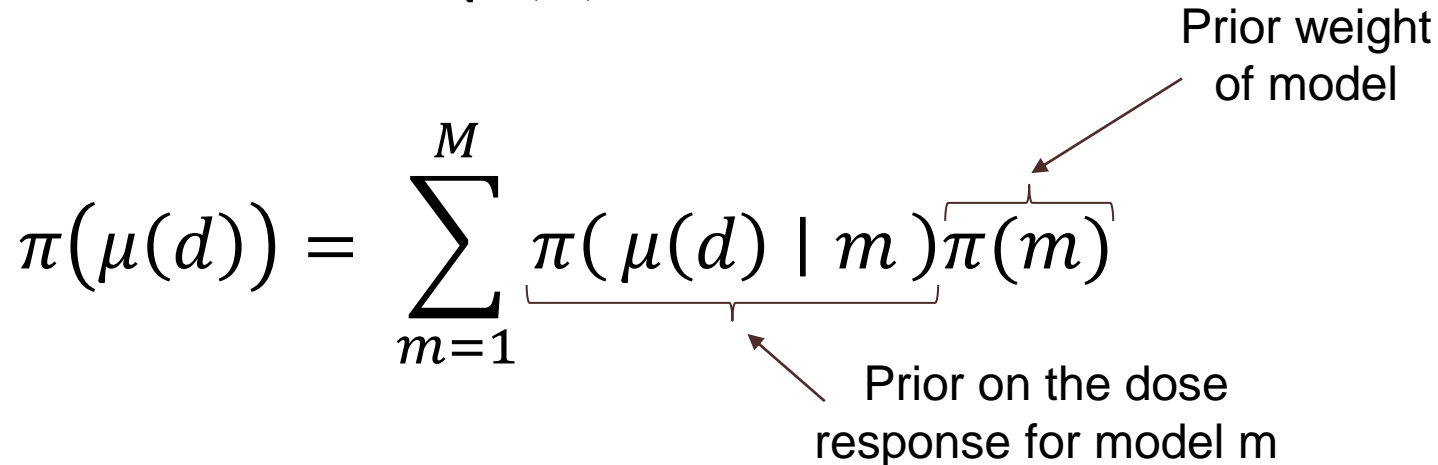
# Bayesian Model Averaging

- Bayesian model averaging is essentially a mixture prior over some quantity of interest.
- In our case, each component of the mixture prior is a different dose-response model.
  - Bayesian analog to the MCP-Mod
  - Bayesian version allows us to include informative priors and/or historical information into the analysis.

# Bayesian Model Averaging

Let  $\mu(d)$  represent the mean response at dose  $d$ .

We can construct a prior over  $M$  different parametric models for  $\mu(d)$ :

$$\pi(\mu(d)) = \sum_{m=1}^M \underbrace{\pi(\mu(d) | m)}_{\text{Prior on the dose response for model } m} \underbrace{\pi(m)}_{\text{Prior weight of model}}$$


# Bayesian Model Averaging

The prior  $\pi(\mu(d) | m)$  for each model is induced from a parametric model.

E.g. Linear Model:  $\mu(d) = \beta_0 + \beta_1 d$

Obtain draws from  $\pi(\beta_0, \beta_1)$  and insert into formula above.

To draw a sample from  $\pi(\mu(d))$  (full Bayesian model averaging prior):

1. Randomly select a model from  $\pi(m)$
2. Randomly draw a set of parameters from that model's prior
3. Obtain  $\mu(d)$  given the parameters drawn in 2.

# Bayesian Model Averaging Posterior

$$p(\mu(d) | y) = \sum_{m=1}^M p(\mu(d) | y, m) p(m | y)$$

Marginal likelihood of data for model  $m$  (integrate over prior)

$$p(m | y) = \frac{\overbrace{p(y | m) \pi(m)}^{\text{Marginal likelihood of data for model } m \text{ (integrate over prior)}}}{\sum_{m^*} p(y | m^*) \pi(m^*)}$$

# Bayesian Model Averaging

The **posterior**  $p(\mu(d) | m, \mathbf{y})$  for each model is induced from a parametric model.

E.g. Linear Model:  $\mu(d) = \beta_0 + \beta_1 d$

Obtain draws from  $p(\beta_0, \beta_1 | \mathbf{y})$  and insert into formula above.

To draw a sample from  $p(\mu(d) | \mathbf{y})$  (Bayesian model averaging **posterior**):

1. Randomly select a model from  $p(m | \mathbf{y})$
2. Randomly draw a set of parameters from that model's posterior
3. Obtain  $\mu(d)$  given the parameters drawn in 2.



# Difficulties of BMA

Calculating the marginals:  $p(y | m)$

- Closed forms usually not available
- Monte Carlo estimates often have high variability and are therefore unreliable.
- One can avoid calculating marginals through one large MCMC chain, e.g., reversible jump MCMC (Green, 1995).
  - Difficult to ensure mixing
  - Computationally intense, often need custom MCMC samplers.

Sensitivity to diffuse priors

# Philosophical Thoughts of BMA

- Do the classical weights using  $p(y | m)$  really give us what we want?
- Essentially  $p(y | m)$  is “how likely is it that this prior generated the parameters which generated the observed data”
- Don't we really want the weights to reflect “which model fits the data best?”

# An Alternative Weight

- Ando & Tsay (2010) replace  $p(y | m)$  with  $\exp(p^*(y | m))$  where  $p^*(y | m)$  is an estimate of the posterior log-predictive likelihood of the observed data for model  $m$ .
- This is justified using a Kullback-Leibler argument (comparing the empirical and posterior predictive distributions).

## Pros

- MCMC can be fit separately for each candidate model.
- Weights are less sensitive to diffuse/non-informative prior choices.
- Calculation of weights can be obtained directly from MCMC output.

## Cons

- Breaks the canonical Bayes' Formula
  - (Is that a bad thing?)
- The estimate of  $p^*(y | m)$  is biased
  - Corrections are suggested by Ando & Tsay (2010), assuming i.i.d data.

# BMA and Dose-Response

- Gould (2019) proposed BMA-Mod which is the Bayesian analog of MCP-Mod.
- Applies Bayesian model averaging with the weights of Ando & Tsay (2010) to dose-response modeling.
- Includes a number of interesting examples.

# Longitudinal Dose-Response Models

*(MANUSCRIPT IN PROGRESS)*

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# Why Longitudinal models?

- Potentially improve decision making earlier
  - Use all available information (e.g. from not-yet-completers)
- Understand the longitudinal response profile for each dose.

# Longitudinal Dose Response

We consider a class of longitudinal dose-response models of the form

$$\mu(d, t) = \alpha + g(d) \times f(t)$$

where  $g(d)$  is a dose-response model,  $f(t)$  is a longitudinal profile at time  $t$ .

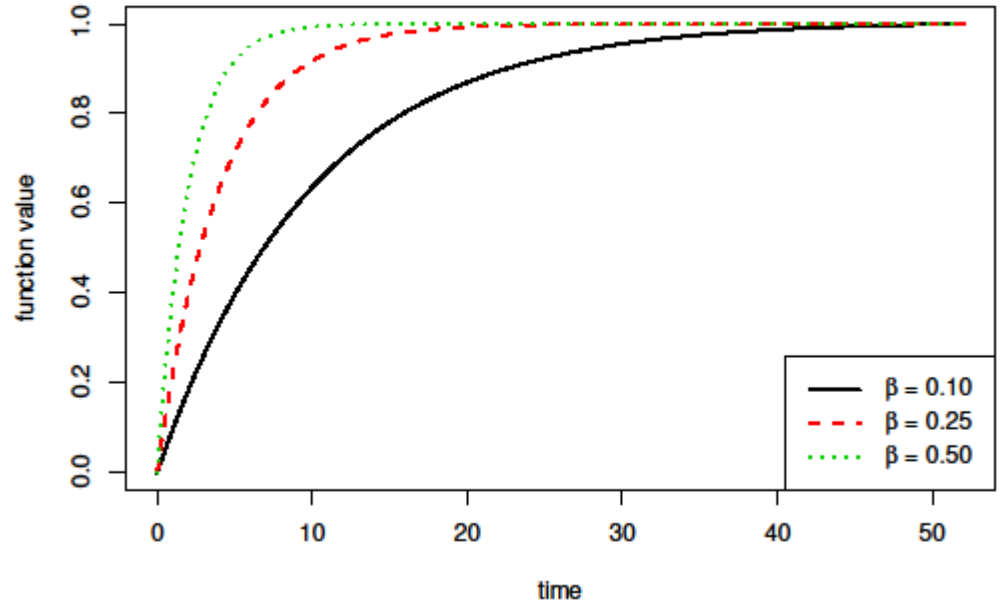
$f(t)$  must satisfy the following conditions:

- Continuous
- $f(0) = 0$
- $0 \leq f(t) \leq 1$
- $\max_t f(t) = 1$

# ITP Model

- Fu and Manner (2010)

$$f(t) = \frac{1 - \exp(-\beta t)}{1 - \exp(-\beta T)}$$



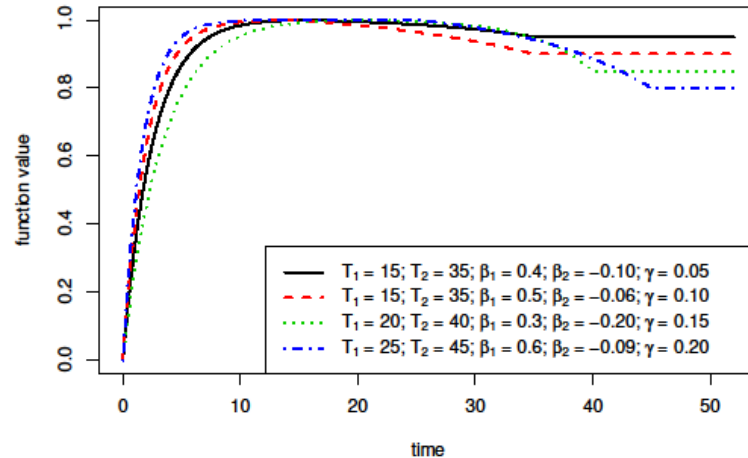


# IDP Model (Pallavi Ray)

$$f(t) = f_1(t)I(0 \leq t < T_1) + f_2(t)I(T_1 \leq t < T_2) + f_2(T_2)I(T_2 \leq t \leq T)$$

$$f_1(t) = \frac{1 - \exp(-\beta_1 t)}{1 - \exp(-\beta_1 T_1)}$$

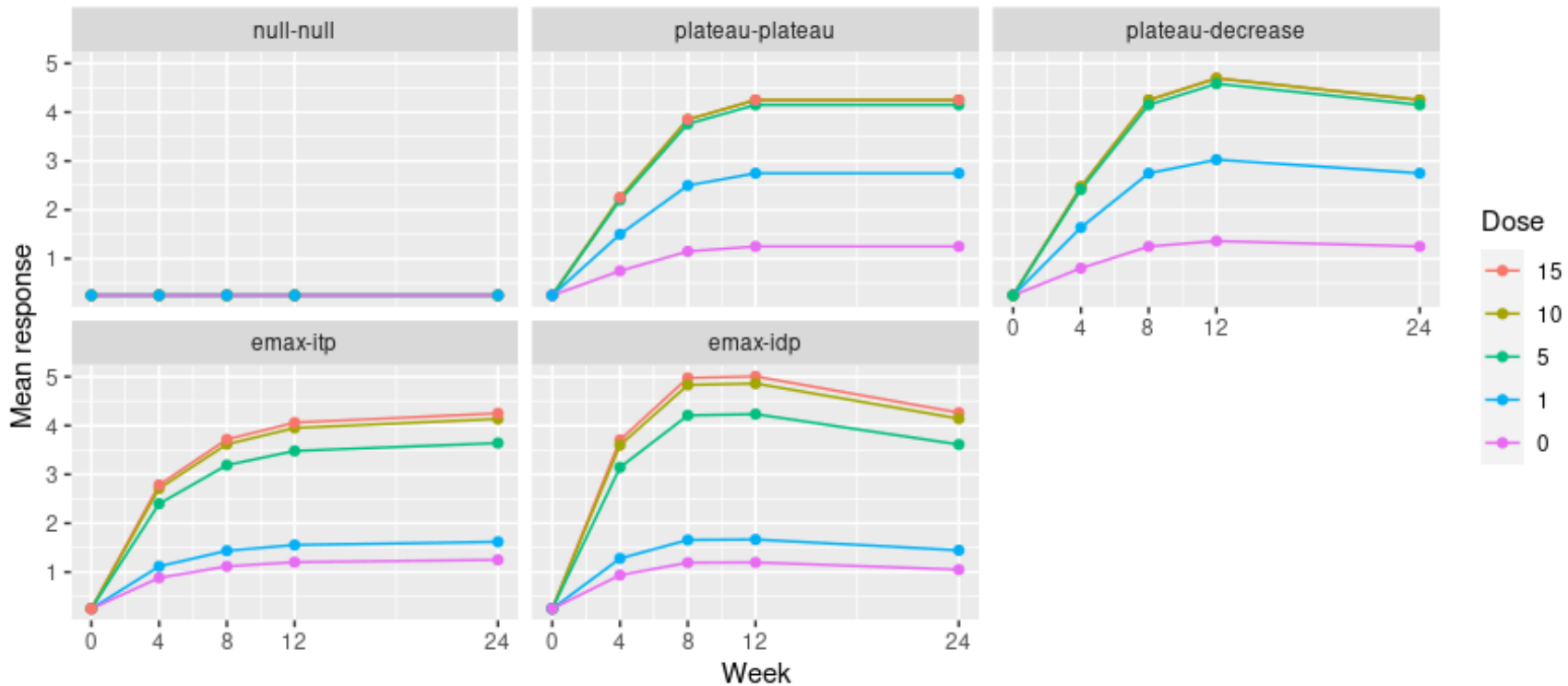
$$f_2(t) = 1 - \gamma \frac{1 - \exp(-\beta_2(t - T_1))}{1 - \exp(-\beta_2(T_2 - T_1))}$$



# Simulations

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# Trial Simulation Scenarios



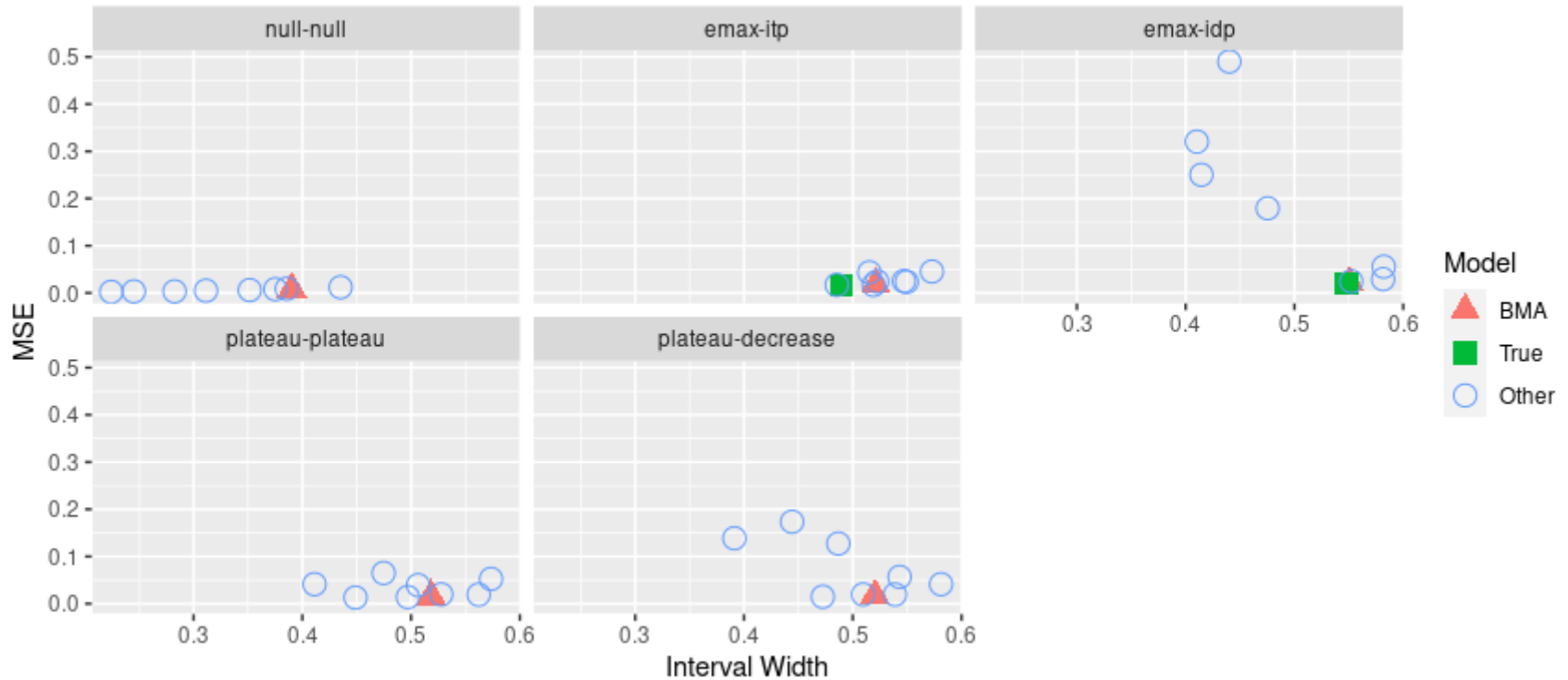
# Setup

- 1:1:1:1:1 randomization over doses
- Interim for early futility or efficacy
  - Interim timing at 50% and 75% completers
- 25 and 50 subjects per arm
- Enrollment of 10 and 30 patients per week
- Analysis: Bayesian model averaging with dose response models: quadratic, log-quadratic, EMAX, exponential crossed with longitudinal ITP and IDP models.
  - 8 total models with equal prior weight.

# MSE



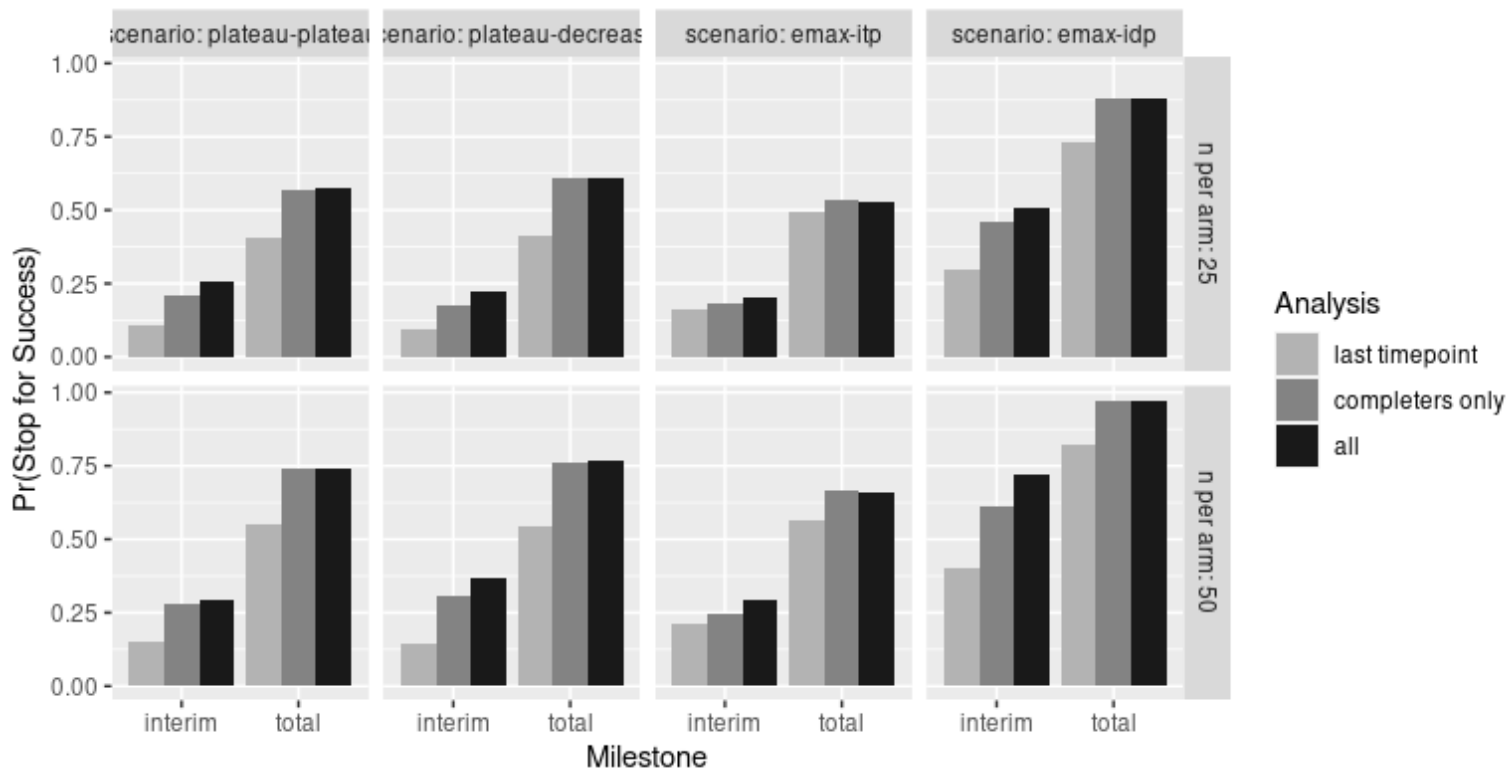
# MSE vs Interval Width



# Posterior Weights



# Interim Stopping





# Future Work

- More complex correlation structures for the longitudinal component.
- Comparison with non-parametric methods

# dreamer R Package

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# dreamer R package

- <https://github.com/rich-payne/dreamer>



# References

- Ando, T., & Tsay, R. (2010). Predictive likelihood for Bayesian model selection and averaging. *International Journal of Forecasting*, 26(4), 744-763.
- Gould, A. L. (2019). BMA-Mod: A Bayesian model averaging strategy for determining dose-response relationships in the presence of model uncertainty. *Biometrical Journal*, 61(5), 1141-1159.
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- Sacks, L. V., Shamsuddin, H. H., Yasinskaya, Y. I., Bouri, K., Lanthier, M. L., & Sherman, R. E. (2014). Scientific and regulatory reasons for delay and denial of FDA approval of initial applications for new drugs, 2000-2012. *Jama*, 311(4), 378-384.